

# PHY604 Lecture 21

November 6, 2025

# Today's lecture:

## Monte Carlo integration and simulation

- Monte Carlo integration with divergences
- Monte Carlo simulation

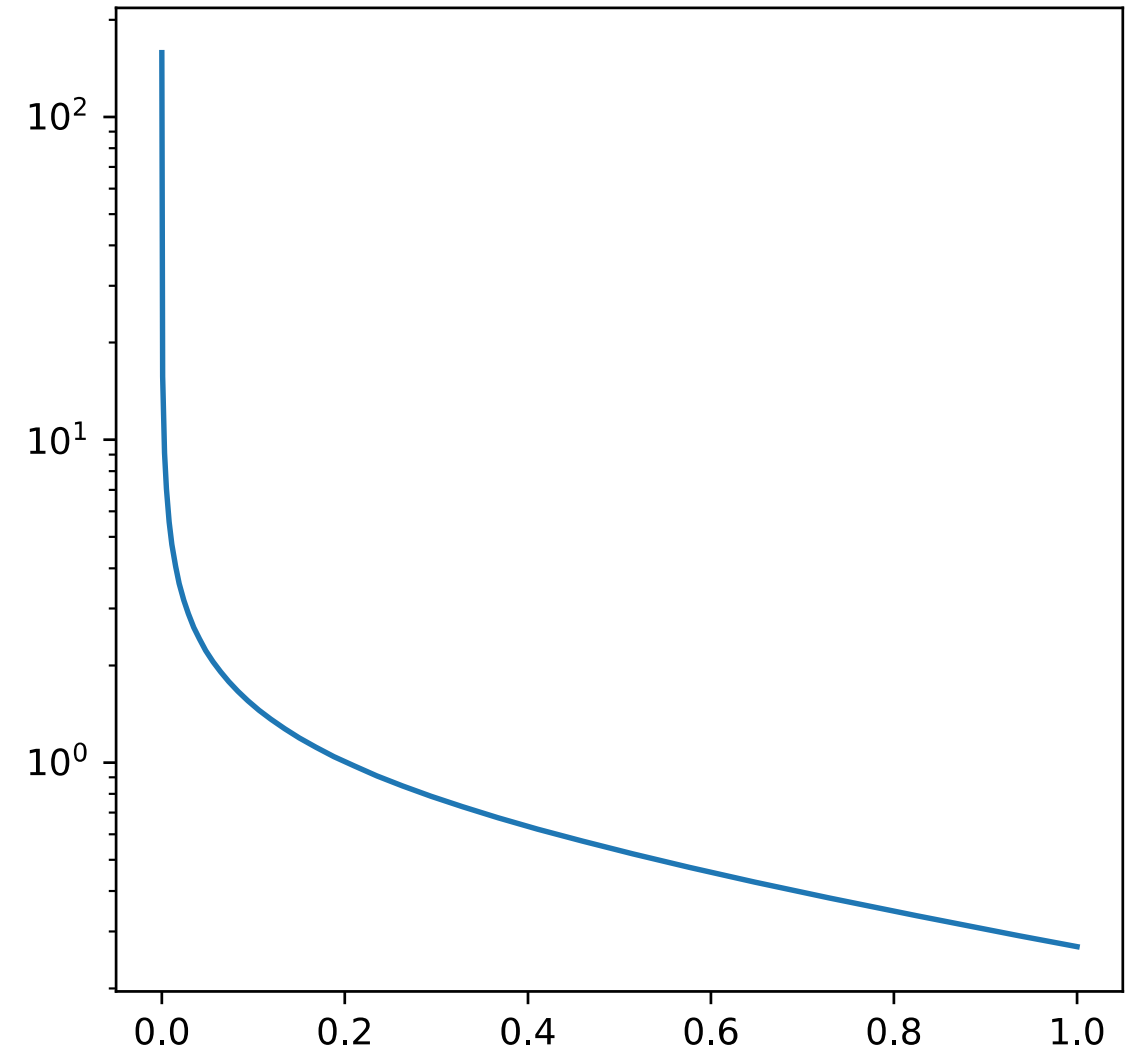
# Monte Carlo integration with divergences

- Monte Carlo integration fails for some pathological functions, e.g., those that contain divergences

- Consider:

$$I = \int_0^1 \frac{x^{-1/2}}{e^x + 1} dx$$

- Function diverges at  $x=0$ , but integral is finite
- E.g., for mean value method, will occasionally get a very large contribution
  - Estimate varies widely between runs



# Importance sampling

- Can get around these issues by drawing points nonuniformly
- For a general function  $g(x)$  can define a **weighted average**:

$$\langle g \rangle_w = \frac{\int_a^b w(x)g(x)dx}{\int_a^b w(x)dx}$$

- $w(x)$  is a weighting function
- If we want to solve a general 1D integral:  $I = \int_a^b f(x)dx$
- We set  $g(x)=f(x)/w(x)$ :

$$\left\langle \frac{f(x)}{w(x)} \right\rangle_w = \frac{\int_a^b f(x)dx}{\int_a^b w(x)dx} = \frac{I}{\int_a^b w(x)dx}$$

# Importance sampling, 1D integral

- Thus, we have:

$$I = \left\langle \frac{f(x)}{w(x)} \right\rangle_w \int_a^b w(x) dx$$

- Equivalent to the mean value method, but from a weighted average
- How do we calculate the weighted average?
- Define probability density function as normalized  $w(x)$

$$p(x) = \frac{w(x)}{\int_a^b w(x) dx}$$

- So

$$\langle g \rangle_w = \int_a^b p(x) g(x) dx$$

# Importance sampling, 1D integral

- Now let's sample  $N$  random points in the interval with the distribution  $p(x)$ . Then:

$$\sum_{i=1}^N g(x_i) \simeq \int_a^b N p(x) g(x) dx$$

- So:

$$\langle g \rangle_w = \int_a^b p(x) g(x) dx \simeq \frac{1}{N} \sum_{i=1}^N g(x_i)$$

- Where  $x_i$  are chosen from the distribution:

$$p(x) = \frac{w(x)}{\int_a^b w(x) dx}$$

# Importance sampling, 1D integral

- Putting everything together:

$$I \simeq \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{w(x_i)} \int_a^b w(x) dx$$

- Generalization of mean value method, which is where  $w(x)=1$
- $w(x)$  can be any function that we choose
  - Can be chosen to remove pathologies in the integrand
- However, now we need to draw from a nonuniform distribution

# Error on importance sampling method

- Error is given by:

$$I_{\text{error}} = \frac{\sqrt{\text{var}_w(f/w)}}{\sqrt{N}} \int_a^b w(x) dx$$

- Where:

$$\text{var}_w g = \langle g^2 \rangle_w - \langle g \rangle_w^2$$

- Still goes like  $N^{-1/2}$

# Importance sampling for pathological function

- Let's return to the integral:  $I = \int_0^1 \frac{x^{-1/2}}{e^x + 1} dx$

- Choose:  $w(x) = x^{-1/2}$

- Then:  $f(x)/w(x) = (e^x + 1)^{-1}$ 
  - Finite and well-behaved over the range

- Probability distribution is:

$$p(x) = \frac{x^{-1/2}}{\int_0^1 x^{-1/2} dx} = \frac{1}{2\sqrt{x}}$$

- So, using the transformation method:

$$\int_0^x \frac{1}{2\sqrt{x'}} dx' = \sqrt{x} = z \quad \implies \quad x = z^2$$

# Importance sampling for pathological function

- So finally, we need to sample:

$$I \simeq \frac{1}{N} \sum_{i=1}^N \frac{f(x_i)}{w(x_i)} \int_a^b w(x) dx = \frac{1}{N} \sum_{i=1}^N \frac{1}{e^{x_i} + 1} \int_0^1 \frac{1}{\sqrt{x}} dx = \frac{1}{N} \sum_{i=1}^N \frac{2}{e^{x_i} + 1}$$

- With the distribution  $x = z^2$

# Today's lecture:

## Monte Carlo integration and simulation

- Monte Carlo integration with divergences
- Monte Carlo simulation

# Monte Carlo simulation

- Any computer simulation that uses random numbers to simulate physical process
- We saw a few examples already: radioactive decay and Rutherford scattering
- Used in every branch of physics
  - Particularly important in **statistical mechanics** and many-body physics

# Monte Carlo simulation in stat mech

- Fundamental problem in statistical mechanics: Calculate expectation value of quantity of interest in thermal equilibrium
- Don't know the exact state of the system, only probability of occupying state  $i$  with energy  $E_i$

$$P(E_i) = \frac{e^{-\beta E_i}}{Z}, \quad Z = \sum_i e^{-\beta E_i}$$

- Then average value of observable  $X$ :

$$\langle X \rangle = \sum_i X_i P(E_i)$$

# States with large numbers

$$\langle X \rangle = \sum_i X_i P(E_i)$$

- Calculating this sum exactly can only be done in a few specific systems (e.g., harmonic oscillator)
- Numerically challenging: states are order Avogadro's number in size
- E.g., one mole of gas with two states: total number of states is  $2^{10^{23}}$
- Instead, use Monte Carlo approach to evaluate the sum

# Monte Carlo approach to expectation values

- We could choose  $N$  terms in the sum at random to add up:

$$\langle X \rangle \simeq \frac{\sum_{k=1}^N X_k P(E_k)}{\sum_{k=1}^N P(E_k)}$$

Needed to normalize the weighted average if not summing over *all* states

- This would not work well! Boltzmann probability is exponentially small for states  $E_i \gg k_B T$
- Usually, most of the states are high energy, only a few contribute significantly
- **Need to use importance sampling!**

# Importance sampling for thermal average

- Choose nonuniform distribution to focus on this small set
- Define weighted average over states:

$$\langle g \rangle_w \simeq \frac{\sum_i w_i g_i}{\sum_i w_i}$$

- We choose:  $g_i = X_i P(E_i) / w_i$

- So:

$$\left\langle \frac{X_i P(E_i)}{w_i} \right\rangle_w = \frac{\sum_i w_i X_i P(E_i) / w_i}{\sum_i w_i} = \frac{\sum_i X_i P(E_i)}{\sum_i w_i} = \frac{\langle X \rangle}{\sum_i w_i}$$

- Or:

$$\langle X \rangle = \left\langle \frac{X_i P(E_i)}{w_i} \right\rangle_w \sum_i w_i$$

# Importance sampling for thermal average

$$\langle X \rangle = \left\langle \frac{X_i P(E_i)}{w_i} \right\rangle_w \sum_i w_i$$

- Evaluate by selecting  $N$  states randomly with nonuniform distribution:

$$\langle X \rangle \simeq \frac{1}{N} \sum_{k=1}^N \frac{X_k P(E_k)}{w_k} \sum_i w_i$$

↑  
Summed over  $N$  samples

← Summed over all states

- Still need to choose  $w_i$  to bias us towards high-probability samples
  - Also, so that sum over all states  $i$  can be evaluated analytically

# Weights for importance sampling

- Simple choice:  $w_i = P(E_i)$
- Sums to 1 over all by definition
- Then we have:

$$\langle X \rangle \simeq \frac{1}{N} \sum_{k=1}^N X_k$$

- Thus, choose  $N$  states in proportion to their Boltzmann weights, and average  $X$  over them

# Markov chain Monte Carlo

- Recall that:  $P(E_i) = \frac{e^{-\beta E_i}}{Z}$ ,  $Z = \sum_i e^{-\beta E_i}$
- Partition function requires a sum over all states that we are trying to avoid
- Can use a **Markov chain** to choose states with probability  $P(E_i)$  without knowing the partition function:
  - Start with a state  $i$
  - Generate a new state  $j$  by making a small change to  $i$
  - Choice of new state is determined probabilistically by a set of **transition probabilities**  $T_{ij}$  that give probability for changing from state  $i$  to  $j$
- If we chose  $T_{ij}$  correctly, probability of visiting any state on a step of the Markov chain is  $P(E_i)$ !

# Transition probabilities in the MC

- We must end up in some state on every MC step, so:

$$\sum_j T_{ij} = 1$$

- Choose transition probabilities such that:

$$\frac{T_{ij}}{T_{ji}} = \frac{P(E_j)}{P(E_i)} = \frac{e^{-\beta E_j} / Z}{e^{-\beta E_i} / Z} = e^{-\beta(E_j - E_i)}$$

- I.e., choosing particular ratio of the probability to go from  $i$  to  $j$ , and  $j$  to  $i$
- Partition function cancels out!

# Transition probabilities in the MC

- If we have correct probability of being in a given state at one step, we will have the correct probability for all later steps

- To see this:

- Suppose we find a set of  $T_{ij}$ 's that satisfy the previous conditions
- Suppose the probability to be in state  $i$  on one particular step is  $P(E_i)$
- Then, probability to be in state  $j$  on the next step is:

$$\sum_i T_{ij} P(E_i) = \sum_i T_{ji} P(E_j) = P(E_j) \sum_i T_{ji} = P(E_j)$$

- Once we get a Boltzmann distribution over states, we will keep it
  - Boltzmann distribution is a fixed point of the Markov chain
- Can also prove that we will converge to Boltzmann distribution
  - See, e.g., Appendix D of Newman

# Metropolis-Hastings accept/reject

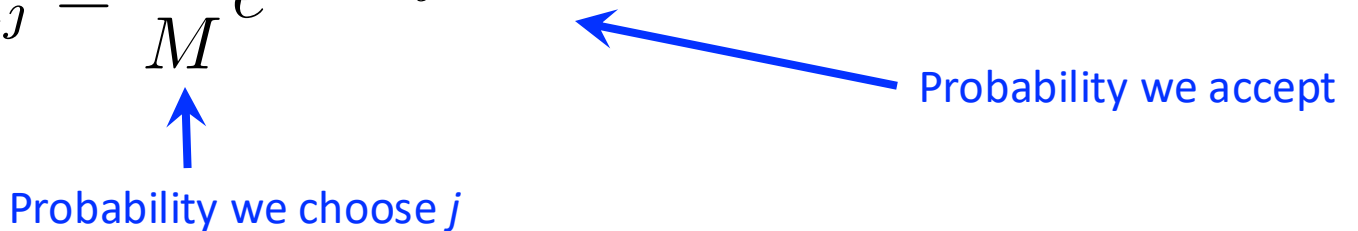
- Still have not worked out what elements of  $T_{ij}$  are
  - Actually, many possible choices
- Most common choice: Metropolis-Hastings algorithm:
  - Choose the change between  $i$  and  $j$  from specified set of possible changes
    - Can be, e.g., chosen at random, uniformly
  - Accept or reject the new state with acceptance probability:

$$P_a = \begin{cases} 1 & \text{if } E_j \leq E_i \\ e^{-\beta(E_j - E_i)} & \text{if } E_j > E_i \end{cases}$$

- I.e., definitely accept if energy is lowered (or equal); may still accept if energy is increased

# Transition probabilities under Metropolis-Hastings

- Total probability to move from  $i$  to given  $j$  (if  $E_j < E_i$ )

$$T_{ij} = \frac{1}{M} e^{-\beta(E_j - E_i)}$$


Probability we choose  $j$

Probability we accept

# Transition probabilities under Metropolis-Hastings

- If  $E_j > E_i$ :

$$T_{ij} = \frac{1}{M} e^{-\beta(E_j - E_i)}, \quad T_{ji} = \frac{1}{M} \quad \implies \quad \frac{T_{ij}}{T_{ji}} = e^{-\beta(E_j - E_i)}$$

- If  $E_j \leq E_i$ :

$$T_{ij} = \frac{1}{M}, \quad T_{ji} = \frac{1}{M} e^{-\beta(E_i - E_j)} \quad \implies \quad \frac{T_{ij}}{T_{ji}} = e^{-\beta(E_j - E_i)}$$

- Thus, both consistent with :

$$\frac{T_{ij}}{T_{ji}} = \frac{P(E_j)}{P(E_i)} = \frac{e^{-\beta E_j} / Z}{e^{-\beta E_i} / Z} = e^{-\beta(E_j - E_i)}$$

# Some comments about the Metropolis algorithm

- Note that many steps will not change the system
  - Still need to include in the sum
- The number of possible moves  $M$ , must be the same when going from  $i$  to  $j$  as  $j$  to  $i$
- Moves must be chosen to get you to every state
  - Move set for which all states are accessible is called **ergodic**
- Will generally take some (unknown) time to equilibrate to Boltzmann distribution

# Steps of Markov chain Monte Carlo:

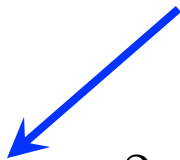
- 1. Choose random starting state
- 2. Choose a move uniformly at random from set of moves
- 3. Calculate the acceptance probability
- 4. Accept or reject the move
- 5. Measure  $X$  in current state, add to sum
- 6. Go back to step 2

# Example: Ideal gas

- Consider the quantum states of a particle or atom of mass  $m$  in cubic box of length  $L$
- Energy of one particle given by:

$$E(n_x, n_y, n_z) = \frac{\pi^2 \hbar^2}{2mL^2} (n_x^2 + n_y^2 + n_z^2)$$

Quantum numbers from  
1 to infinity.



- Ideal gas: no interactions between particles
  - Energy is sum of individual particles:

$$E = \sum_{i=1}^N E(n_x^{(i)}, n_y^{(i)}, n_z^{(i)})$$

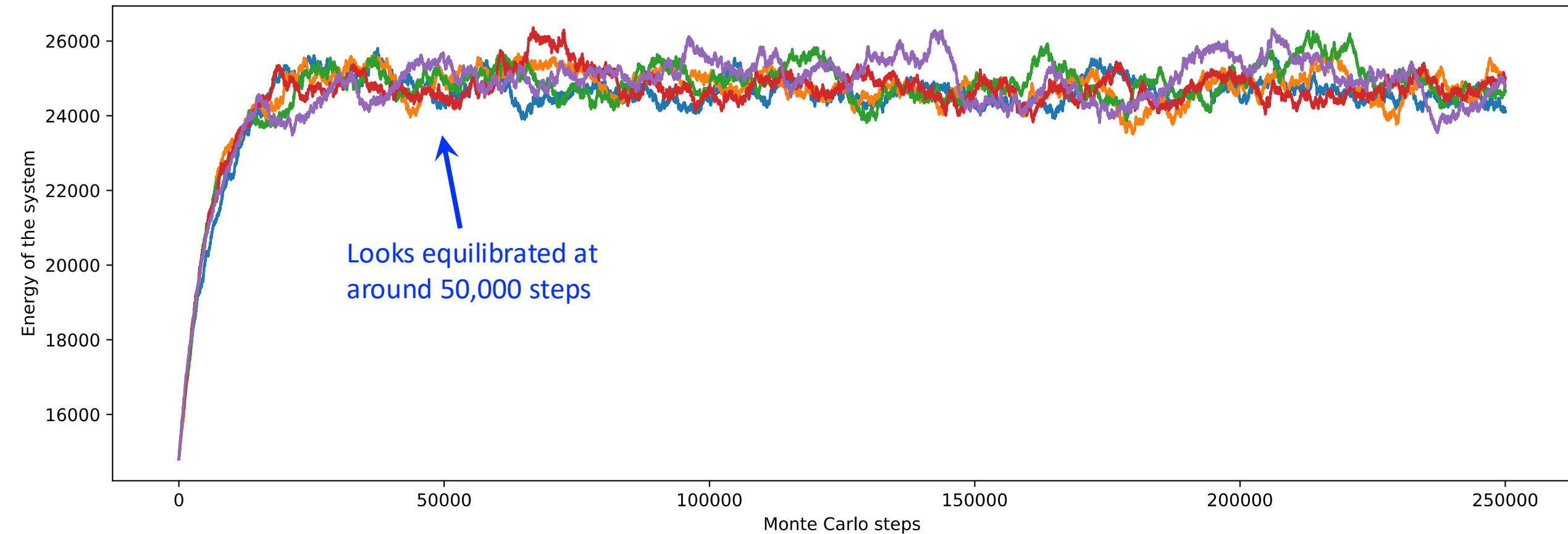
# Move set for ideal gas

- Choose set of all moves of a single atom to one of the six “neighboring” states where  $n_x$ ,  $n_y$ , or  $n_z$  differ by  $\pm 1$
- Each Monte Carlo step, choose a random particle, choose a quantum number, change it by  $\pm 1$
- Change in total energy just the change for single particle since there are no interactions
  - E.g., increase or decrease  $n_x$  of atom  $i$  by one:

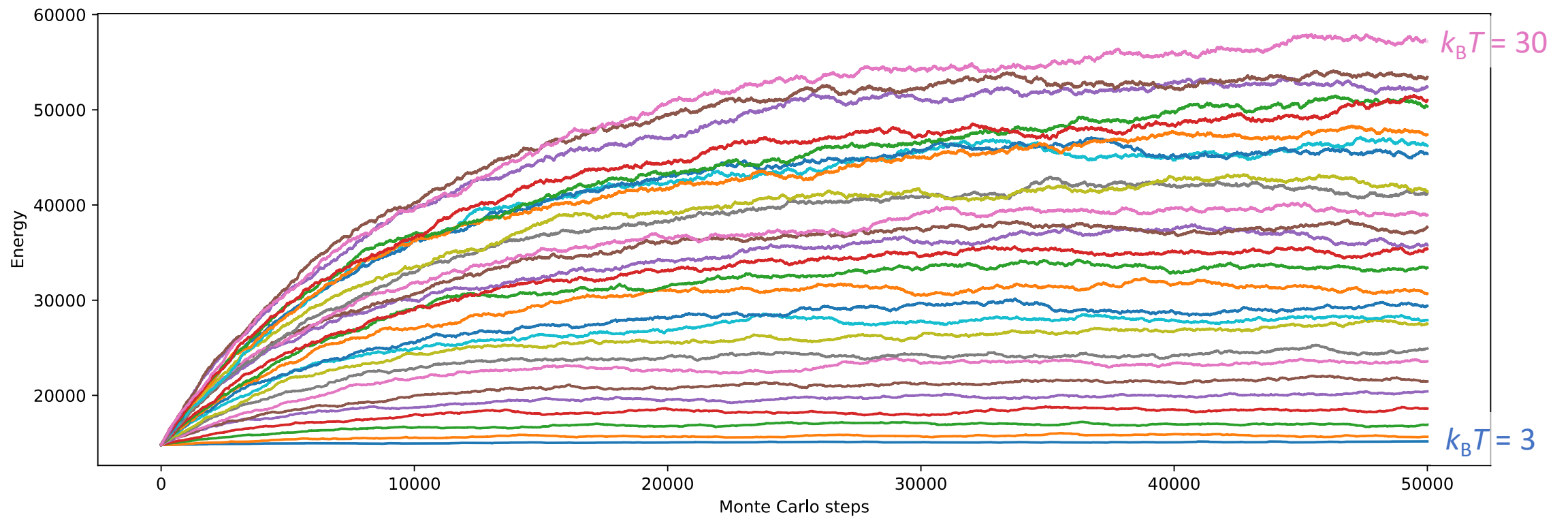
$$\begin{aligned}\Delta E &= \frac{\pi^2 \hbar^2}{2mL^2} [(n_x \pm 1)^2 + n_y^2 + n_z^2] - \frac{\pi^2 \hbar^2}{2mL^2} (n_x^2 + n_y^2 + n_z^2) \\ &= \frac{\pi^2 \hbar^2}{2mL^2} [(n_x \pm 1)^2 - n_x^2] = \frac{\pi^2 \hbar^2}{2mL^2} (\pm 2n_x + 1)\end{aligned}$$

- Note: Reject moves that try to make  $n < 1$

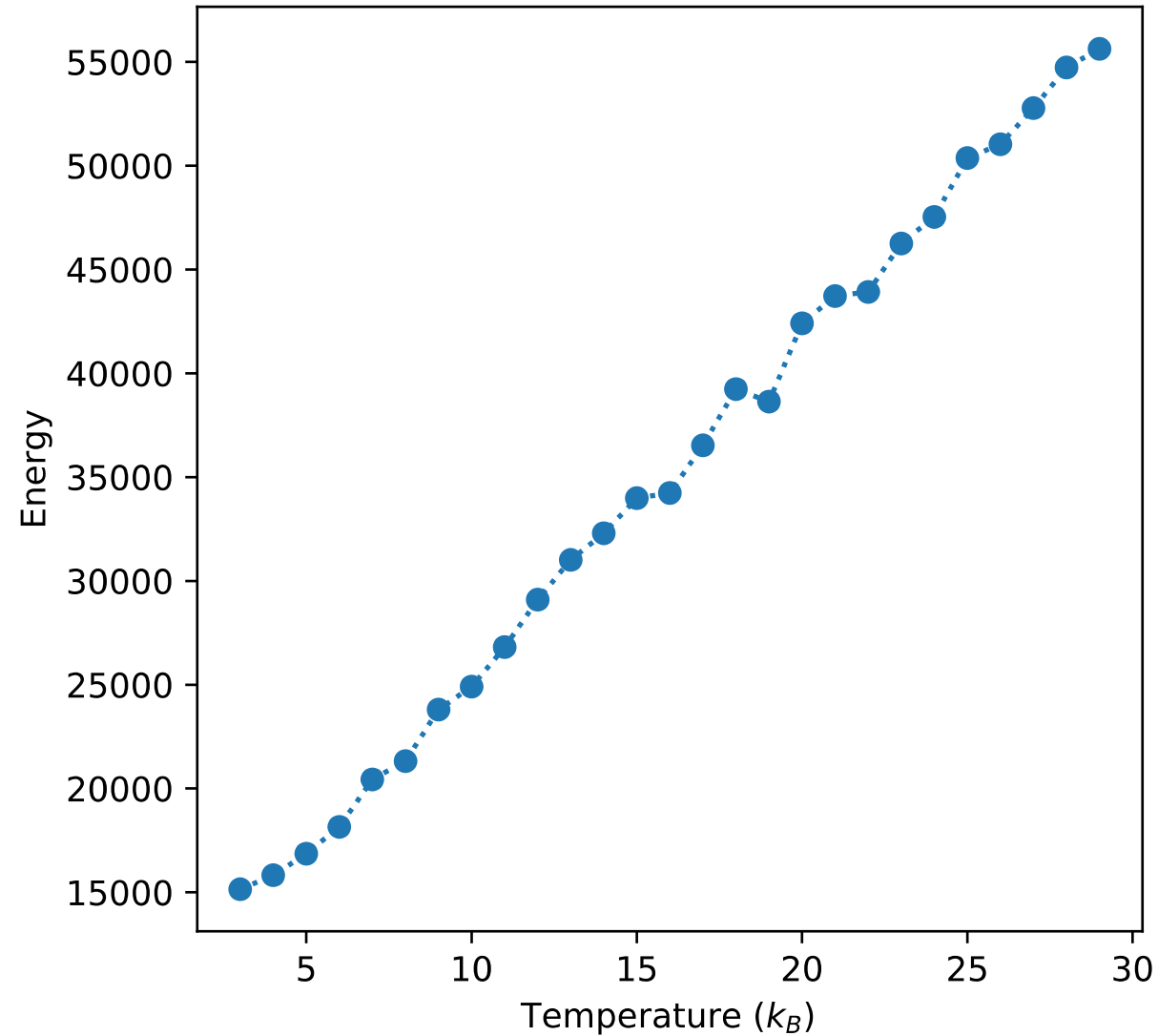
# Monte Carlo simulation of ideal gas



# Monte Carlo simulation of ideal gas: Dependence on T



# Monte Carlo simulation of ideal gas: E vs. T



# After class tasks

- Homework 5 is posted, due Nov. 19, 2025
- Almost done grading Homework 3
- Final project ideas due Nov. 19
- Readings:
  - Newman Sec. 10.3